HTAP2 multi-model estimates of premature human mortality due to intercontinental transport of air pollution and emission sectors

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Abstract. Ambient air pollution from ozone and fine particulate matter is associated with premature mortality. As emissions from one continent influence air quality over others, changes in emissions can also influence human health on other continents. We estimate global air-pollution-related premature mortality from exposure to PM$_{2.5}$ and ozone and the avoided deaths due to 20% anthropogenic emission reductions from six source regions, North America (NAM), Europe (EUR), South Asia (SAS), East Asia (EAS), Russia–Belarus–Ukraine (RBU), and the Middle East (MDE), three global emission sectors, power and industry (PIN), ground transportation (TRN), and residential (RES), and one global domain (GLO), using an ensemble of global chemical transport model simulations coordinated by the second phase of the Task Force on Hemispheric Transport of Air Pollutants (TF HTAP2), and epidemiologically derived concentra-
tion response functions. We build on results from previous studies of TF HTAP by using improved atmospheric models driven by new estimates of 2010 anthropogenic emissions (excluding methane), with more source and receptor regions, new consideration of source sector impacts, and new epidemiological mortality functions. We estimate 290,000 (95% confidence interval (CI): 30,000, 600,000) premature O$_3$-related deaths and 2.8 million (0.5 million, 4.6 million) PM$_{2.5}$-related premature deaths globally for the baseline year 2010. While 20% emission reductions from one region generally lead to more avoided deaths within the source region than outside, reducing emissions from MDE and RBU can avoid more O$_3$-related deaths outside of these regions than within, and reducing MDE emissions also avoids more PM$_{2.5}$-related deaths outside of MDE than within. Our findings that most avoided O$_3$-related deaths from emission reductions in NAM and EUR occur outside of those regions contrast with those of previous studies, while estimates of PM$_{2.5}$-related deaths from NAM, EUR, SAS, and EAS emission reductions agree well. In addition, EUR, MDE, and RBU have more avoided O$_3$-related deaths from reducing foreign emissions than from domestic reductions. For six regional emission reductions, the total avoided extra-regional mortality is estimated as 6000 (−3400, 15,500) deaths per year and 25,100 (8200, 35,800) deaths per year through changes in O$_3$ and PM$_{2.5}$, respectively. Interregional transport of air pollutants leads to more deaths through changes in PM$_{2.5}$ than in O$_3$, even though O$_3$ is transported more on interregional scales, since PM$_{2.5}$ has a stronger influence on mortality. For NAM and EUR, our estimates of avoided mortality from regional and extra-regional emission reductions are comparable to those estimated by regional models for these same experiments. In sectoral emission reductions, TRN emissions account for the greatest fraction (26–53% of global emission reduction) of O$_3$-related premature deaths in most regions, in agreement with previous studies, except for EAS (58%) and RBU (38%) where PIN emissions dominate. In contrast, PIN emission reductions have the greatest fraction (38–78% of global emission reduction) of PM$_{2.5}$-related deaths in most regions, except for SAS (45%) where RES emission dominates, which differs with previous studies in which RES emissions dominate global health impacts. The spread of air pollutant concentration changes across models contributes most to the overall uncertainty in estimated avoided deaths, highlighting the uncertainty in results based on a single model. Despite uncertainties, the health benefits of reduced intercontinental air pollution transport suggest that international cooperation may be desirable to mitigate pollution transported over long distances.

1 Introduction

Ozone (O$_3$) and fine particulate matter with an aerodynamic diameter of less than 2.5 µm (PM$_{2.5}$) are two common air pollutants with known adverse health effects. Epidemiological studies have shown that both short-term and long-term exposures to O$_3$ and PM$_{2.5}$ are associated with elevated rates of premature mortality. Short-term exposure to O$_3$ is associated with respiratory morbidity and mortality (Bell et al., 2005, 2014; Gryparis et al., 2004; Ito et al., 2005; Levy et al., 2005; Stieb et al., 2009) while long-term exposure to O$_3$ has been associated with premature respiratory mortality (Jerrett et al., 2009; Turner et al., 2016). Short-term exposure to PM$_{2.5}$ has been associated with increases in daily mortality rates from all natural causes, and specifically from respiratory and cardiovascular causes (Bell et al., 2014; Du et al., 2016; Powell et al., 2015; Pope et al., 2011), while long-term exposure to PM$_{2.5}$ can have detrimental chronic health effects, including premature mortality due to cardiopulmonary diseases and lung cancer (Brook et al., 2010; Burnett et al., 2014; Hamra et al., 2014; Krewski et al., 2009; Lepeule et al., 2012; Lim et al., 2012). The Global Burden of Disease Study 2015 (GBD 2015) estimated 254,000 deaths per year associated with ambient O$_3$ and 4.2 million associated with ambient PM$_{2.5}$ (Cohen et al., 2017). A comparable study using output from an ensemble of global chemistry–climate models estimated 470,000 deaths per year associated with O$_3$ and 2.1 million premature deaths per year associated with anthropogenic PM$_{2.5}$ (Silva et al., 2013). These differences in GBD estimates result mainly from differences in concentration response functions and estimates of pollutant concentrations.

Numerous observational and modeling studies have shown that anthropogenic emissions can affect O$_3$ and PM$_{2.5}$ concentrations across continents (Dentener et al., 2010; Heald et al., 2006; Leibensperger et al., 2011; Lin et al., 2012, 2017; Liu et al., 2009a; West et al., 2009a; Wild and Aki-moto, 2001; Yu et al., 2008). As changes in emissions from one continent influence air quality over others, several studies have estimated the premature mortality from intercontinental transport (Anenberg et al., 2009, 2014; Bhalla et al., 2014; Duncan et al., 2008; Im et al., 2018; Liu et al., 2009b; West et al., 2009b; Zhang et al., 2017). In 2005, the Task Force on Hemispheric Transport of Air Pollutants (TF HTAP) was launched under the United Nations Economic Commission for Europe (UNECE) Convention on Long-Range Transboundary Air Pollution (LRTAP). One of its tasks is to investigate the impacts of emission reductions on the intercontinental transport of air pollution, air quality, health, ecosystem, and climate effects using a multi-model ensemble to quantify uncertainties due to differences among models (Anenberg et al., 2009, 2014; Fiore et al., 2009; Fry et al., 2012; Huang et al., 2017; Stjern et al., 2016; Yu et al., 2013).

In TF HTAP Phase 1 (TF HTAP1), human premature mortality due to 20% anthropogenic emission reductions in four large source regions was investigated by Anenberg et al. (2009, 2014). They found that 20% foreign O$_3$ precursor emission reductions contribute approximately 30 to > 50% of the deaths avoided by reducing precursor emis-
In this study, we estimate the impacts of interregional transport and of source sector emissions on human premature mortality from PM$_{2.5}$, using an ensemble of global chemical transport models coordinated by the Task Force on Hemispheric Transport of Air Pollutants Phase 2 (TF HTAP2) (Galmarini et al., 2017; Huang et al., 2017; Janssens-Maenhout et al., 2015; Stjern et al., 2016). Anthropogenic emissions were reduced by 20% in six source regions, North America (NAM), Europe (EUR), South Asia (SAS), East Asia (EAS), Russia–Belarus–Ukraine (RBU), and the Middle East (MDE), three emission sectors, power and industry (PIN), ground transportation (TRN), and residential (RES), and one worldwide region (GLO). Human premature mortality due to these reductions is calculated using a health impact function based on a log-linear model for PM$_{2.5}$ (Jerrett et al., 2009) and an integrated exposure–response (IER) model for PM$_{2.5}$ (Burnett et al., 2014), within the six source regions and elsewhere in the world. We conduct a Monte Carlo simulation to estimate the overall uncertainty due to uncertainties in relative risk, air pollutant concentrations (given by the spread of results among different models), and baseline mortality rates.

2 Methods

2.1 Modeled O$_3$ and PM$_{2.5}$ surface concentration

Global numerical modeling experiments initiated by TF HTAP2, the regional experiments by the Air Quality Model Evaluation International Initiative (AQMEII) over EUR and NAM, and the Model Intercomparison Study-Asia (MICS-Asia) were coordinated to perform consistent emission perturbation modeling experiments across the global, hemispheric, and continental/regional scales (Galmarini et al., 2017). Simulation periods, meteorology, emission inventories, boundary conditions, and model output are also consistent. The Joint Research Centre’s (JRC) EDGAR (Emission Data Base for Global Research) team in collaboration with regional emission experts from the U.S. Environmental Protection Agency (US EPA), European Monitoring and Evaluation Programme (EMEP), Centre on Emission Inventories and Projections (CEIP), Netherlands Organization for Applied Scientific Research (TNO), and the MICS-Asia Scientific Community and Regional Emission Activity Asia (REAS) provide a global emission inventory at 0.1° × 0.1° resolution for TF HTAP2 modeling experiments (Janssens-Maenhout et al., 2015). The emission dataset was constructed for SO$_2$, NO$_x$, CO, non-methane volatile organic compounds, NH$_3$, PM$_{10}$, PM$_{2.5}$, black carbon (BC), and organic carbon (OC) and seven emission sectors (shipping, aircraft, land transportation, agriculture, residential, industry, and energy) for the year 2010 (Fig. S1 in the Supplement). This study uses outputs from 14 global models/model versions (Table S1) participating in TF HTAP2. Overall, TF HTAP2 model resolutions are finer than in TF HTAP1. In TF HTAP2, each model performed a baseline simulation.
and sensitivity simulations in which the anthropogenic emissions in a defined source region or sector were perturbed (reduced by 20% in most cases). Based on the number of models that simulated different experiments, we choose to focus on emission reductions from six source regions, three emission sectors, and one global domain. More specifically, all anthropogenic emissions are reduced by 20% in the NAM, EUR, SAS, EAS, RBU, and MDE continental regions, in the PIN, TRN, and RES emission sectors globally, and in GLO (Fig. S2). Unlike TF HTAP1 (Dentener et al., 2010), which defined rectangular regions that included ocean or some sparsely inhabited regions, TF HTAP2 regions are defined by geopolitical boundaries.

We selected output from the models that provided temporally resolved volume mixing ratios of O$_3$ and mass mixing ratios of PM$_{2.5}$ ("mmrpm2p5") for the baseline and at least one regional or sectoral emission reduction scenario. Among the 14 models, 11 models reported O$_3$ and eight reported PM$_{2.5}$ for regional emission perturbation scenarios. Four models reported O$_3$ and four reported PM$_{2.5}$ for sectoral emission perturbation scenarios, and 10 models reported O$_3$ and eight reported PM$_{2.5}$ for the global emission perturbation. All models used prescribed meteorology for the year 2010, although this meteorology was derived from different (re)analysis products and was not uniform across models. Modeled concentrations are processed by calculating metrics consistent with the underlying epidemiological studies to estimate premature mortality. For O$_3$, we calculate the average of daily 1 h maximum O$_3$ concentration for the 6 consecutive months with the highest concentrations in each grid cell (Jerrett et al., 2009), for the baseline and each 20% emission reduction scenario. While some models reported hourly O$_3$ metrics, others only reported daily or monthly O$_3$. We include these models by first calculating the ratio of the 6-month average of daily 1 h maximum O$_3$ to the annual average of O$_3$ in individual grid cells, for models reporting hourly O$_3$, and then applying that ratio to the annual average of ozone for those models that only report daily or monthly O$_3$, following Silva et al. (2013, 2016b). For PM$_{2.5}$, we calculate the annual average PM$_{2.5}$ concentration in each cell using the monthly total PM$_{2.5}$ concentrations reported by each model ("mmrpm2p5"). Model results for these two metrics are then regridded from each model’s native grid resolution (varying from 0.5° × 0.5° to 2.8° × 2.8°) to a consistent 0.5° × 0.5° resolution used in mortality estimation. We estimate regional and sectoral multi-model averages for each 20% emission reduction scenario in the year 2010, but for each perturbation case, we only include models that report both the baseline and perturbation cases.

2.2 Model evaluation

Measurements from multiple observation networks are employed in this study to evaluate the model performance around the world. We evaluate model performance for the 2010 baseline simulation for 11 TF HTAP2 models for O$_3$ and eight models for PM$_{2.5}$ (Table S1). For O$_3$, we use ground level measurements from 2010 at 4655 sites globally, collected by the Tropospheric Ozone Assessment Report (TOAR) (Schultz et al., 2017; Young et al., 2018). The TOAR dataset identifies stations as urban, rural, and unclassified sites (Schultz et al., 2017). Model performance is evaluated for the average of daily 1 h maximum O$_3$ concentrations for the 3 consecutive months (3m1hmaxO$_3$) with the highest concentrations in each grid cell, including models that only report daily or monthly O$_3$ as described above. This metric for O$_3$ differs slightly from the 6-month average of daily 1 h maximum metric used for health impact assessment and is chosen because TOAR reports the 3-month metric but not the 6-month metric. For PM$_{2.5}$, we compare the annual average PM$_{2.5}$, using PM$_{2.5}$ observations from 2010 at 3157 sites globally selected for analysis by the Global Burden of Disease 2013 (GBD2013) (Forouzanfar et al., 2016; Brauer et al., 2016). Statistical parameters including the normalized mean bias (NMB), normalized mean error (NME), and correlation coefficient (R) are selected to evaluate model performance.

Tables S2 and S3 present statistical parameters of model evaluation for O$_3$ and PM$_{2.5}$, and Figs. S3–S10 show the spatial O$_3$ and PM$_{2.5}$ evaluation as NMB around the world and in NAM, EUR, and EAS. For 3m1hmaxO$_3$, the model ensemble mean shows good agreement with measurements globally with a NMB of 7.3% and NME of 13.2% but moderate correlation with R of 0.53 (Table S2). For individual models, eight models (CAM-Chem, CHASER_T42, CHASER_T106, EMEPrv48, GEOS-Chem Adjoint, GEOS-Chem, GFDL_AM3, and HadGEM2-ES) overestimate 3m1hmaxO$_3$ with NMB of 9.2 to 23% while three models (C-IFS, OsloCTM3_v2, and RAQMS) underestimate 3m1hmaxO$_3$ by −10.8 to −19.4% globally (Fig. S3). In the six perturbation regions, the model ensemble mean is also in good agreement with the measurements, with −11.2 to 25.3% for NMB, 9.8 to 25.3% for NME, and −0.09 to 0.98 for R. The ranges of NMB for individual models are −18.1 to 32.3%, −24.1 to 21.3%, −24.5 to 45.0%, −26.4 to 24.5%, −30.5 to 20.3%, and −35.3 to 5.4% in NAM, EUR, SAS, EAS, MDE, and RBU, respectively (Figs. S4–S6). Note that some regions (SAS, MDE, and RBU) have very few observations for model evaluation, making the comparison less robust. The underestimated O$_3$ in the western US and overestimated O$_3$ in the eastern US in most models are very close to the model performance results of Huang et al. (2017), who compare eight TF HTAP2 models with CASTNET observations (Fig. S4), as well as earlier studies under HTAP1 (Fiore et al., 2009). Similarly, Dong et al. (2018) find that O$_3$ is overestimated in EUR and EAS by six TF HTAP2 models, consistent with our ensemble mean result in these two regions (Figs. S5–S6).

For PM$_{2.5}$, the model ensemble mean agrees well with measurements globally, with a NMB of −23.1%, NME of
35.4 %, and \( R = 0.77 \) (Table S3). For individual models, only one model (GEOS-Chem Adjoint) overpredicts PM\(_{2.5}\) by 20.3 %, while the other seven models underpredict PM\(_{2.5}\) by \(-60.9\) to \(-7.4\) % around the world (Fig. S7). In six perturbation regions, the model ensemble mean is also in good agreement with measurements, with ranges of NMB of \(-49.7\) to \(-1.4\) %, \(21.2\) to \(-49.7\) % for NNE, and \(-0.5\) to \(-1.0\) for \(R\). The ranges of NMB for individual models are \(-46.6\) to \(-13.9\), \(-76.0\) to \(-31.9\), \(-35.0\) to \(-49.7\), \(-50.4\) to \(-29.5\), \(-52.6\) to \(-31.5\), and \(-74.1\) to \(-19.8\) % in NAM, EUR, SAS, EAS, MDE, and RBU, respectively (Figs. S8–S10). Dong et al. (2018) show that PM\(_{2.5}\) is underestimated in EUR and EAS by six TF HTAP2 models, consistent with our ensemble mean result in these two regions (Figs. S9–S10). Note that many observations used are located in urban areas, and models with coarse resolution may not be expected to have good model performance. Several models also neglect some PM\(_{2.5}\) species, which may explain the tendency of models to underestimate.

### 2.3 Health impact assessment

We use output from the TF THAP2 model ensemble to estimate annual global O\(_3\) and PM\(_{2.5}\)-related cause-specific premature mortality and avoided mortality from the 20 % regional and sectoral emission reductions, following the same methods used by Silva et al. (2016a, b). The annual O\(_3\)- and PM\(_{2.5}\)-related premature mortality is calculated using a health impact function based on epidemiological relationships between ambient air pollution concentration and mortality in each grid cell: \(\Delta M = \gamma_0 \times AF \times Pop\), where \(\Delta M\) is premature mortality, \(\gamma_0\) is the baseline mortality rate (for the exposed population), \(AF = 1 - 1/RR\) is the attributable fraction, where \(RR\) is relative risk of death attributable to the change in air pollutant concentration (\(RR = 1\) when there is no increased risk of death associated with a change in pollutant concentration), and Pop is the exposed population (adults aged 25 and older).

For O\(_3\) mortality, we use a log-linear model for chronic respiratory mortality (RESP) from an American Cancer Society (ACS) study (Jerrett et al. 2009), following recent studies including the GBD (Cohen et al., 2017), but Turner et al. (2016) recently published new results for chronic ozone mortality, and adoption of these results would lead to more ozone-related deaths overall (Malley et al., 2017). \(RR\) is calculated as

\[
RR = e^{\beta \Delta x}, \tag{1}
\]

where \(\beta\) is the concentration response factor, and \(\Delta x\) corresponds to the change in pollutant concentrations among simulations with perturbed emissions and the baseline simulation. For O\(_3\), \(RR = 1.040\) (95 % CI: 1.013–1.067) for a 10 ppb increase in O\(_3\) concentrations (Jerrett et al., 2009), which from Eq. (1) gives values for \(\beta\) of 0.00392 (0.00129–0.00649). We estimate O\(_3\)-related premature deaths due to respiratory disease (RESP) based on decreases or increases in O\(_3\) concentration (i.e., \(\Delta x\)) due to 20 % regional and sectoral emission reduction scenarios relative to the baseline. For regional and sectoral reductions, we do not assume a low-concentration threshold below which changes in O\(_3\) have no mortality effects as there is no clear evidence for such a threshold, following Anenberg et al. (2009, 2010) and Silva et al. (2013, 2016a, b). However, we evaluate global O\(_3\) premature mortality for the baseline 2010 simulation, relative to a counterfactual concentration of 37.6 ppb (Lim et al., 2012), for consistency with GBD estimates (Cohen et al., 2017).

For PM\(_{2.5}\) mortality, we apply the IER model, which is intended to better represent the risk of exposure to PM\(_{2.5}\) at locations with high ambient concentrations (Burnett et al., 2014). \(RR\) is calculated as

\[
\text{For } z < z_{cf}, \quad RR_{IER} (z) = 1, \tag{2}
\]

\[
\text{For } z \geq z_{cf}, \quad RR_{IER} (z) = 1 + \alpha [1 - \exp(-\gamma (z - z_{cf})^\delta)], \tag{3}
\]

where \(z\) is the PM\(_{2.5}\) concentration in micrograms per cubic meter and \(z_{cf}\) is the counterfactual concentration below which no additional risk is assumed, and the parameters \(\alpha\), \(\gamma\), and \(\delta\) are used to fit the function for cause-specific RR (Burnett et al., 2014). The overall PM\(_{2.5}\)-related cause-specific premature deaths related to ischemic heart disease (IHD), cerebrovascular disease (STROKE), chronic obstructive pulmonary disease (COPD), and lung cancer (LC) are estimated using RRs per age group for IHD and STROKE and RRs for all ages for COPD and LC. A uniform distribution from 5.8 to 8.8 \(\mu g m^{-3}\) is used for \(z_{cf}\) as suggested by Burnett et al. (2014), which does not vary in space nor time. For uncertainty analysis, we use results from 1000 Monte Carlo simulations of Burnett et al. (2014) to calculate RR in each grid cell using Eq. (2) or (3). We estimate avoided premature mortality in 20 % emission perturbation experiments by taking the difference in premature mortality estimates with the 2010 baseline. However, in the IER model, the concentration response function flattens off at higher PM\(_{2.5}\) concentrations, yielding different estimates of avoided premature mortality for identical changes in air pollutant concentrations from less-polluted vs. highly polluted regions. That is, one unit reduction of air pollution may have a stronger effect on avoided mortality in regions where pollution levels are lower (e.g., EUR, NAM) compared with highly polluted regions (e.g., EAS, India), which would not be the case for a log-linear function (Jerrett et al., 2009; Krewski et al., 2009). Therefore, using the IER model in this study may result in smaller changes in avoided mortality in highly polluted areas than using the linear model.

For the exposed population, we use the Oak Ridge National Laboratory’s LandScan 2011 Global Population Database at approximately 1 km resolution (30° × 30°) (Bright et al., 2012). For the population of adults aged 25 and older, we use ArcGIS 10.2 geoprocessing tools to es-
timate the population per 5-year age group in each cell by multiplying the country level percentage in each age group by the population in each cell. We obtained cause-specific baseline mortality rates for 187 countries from the GBD 2010 mortality dataset (IHME, 2013). The population and baseline mortality per age group were regridded to the 0.5° × 0.5° grid (Table S4 and Fig. S11). Cause-specific baseline mortality rates vary geographically, e.g., RESP and COPD are relatively more dominant in SAS, IHD in EUR, STROKE in Russia, and LC in NAM.

Finally, we conduct 1000 Monte Carlo simulations to propagate uncertainty from baseline mortality rates, modeled air pollutant concentrations, and the RRs in health impact functions. We use the reported 95% CIs for cause-specific baseline mortality rates, assuming lognormal distributions. For modeled O$_3$ and PM$_{2.5}$ concentrations we use the absolute value of the coefficient of variation among models in each grid cell, for each 20% emission perturbation case minus the baseline, assuming a normal distribution. For O$_3$ RRs, we use the reported 95% confidence intervals (CIs), assuming a normal distribution. For PM$_{2.5}$ RRs, we use the parameter values (i.e., $\alpha$, $\gamma$, $\delta$, and $\varepsilon$) of Burnett et al. (2014) for 1000 simulations. One should acknowledge that the range of modeled air pollution concentrations in an ensemble is not a true reflection of the uncertainty in emissions or concentrations. The mean health outcome of the 1000 Monte Carlo simulations (the empirical mean) may differ from the mean when using the mean RR.

We also quantify the uncertainties in mortality due to the spread of air pollutant concentrations across models, RRs, and baseline mortality rates, as contributors to the overall uncertainty, expressed as a coefficient, of variation and compare the result with the Monte Carlo analysis estimate. To do so, we hold two variables at their mean values and change the variable of interest within its uncertainty range; for example, using mean RRs and baseline mortality rates, we analyze the spread of the model ensemble to calculate the coefficient of variation caused by model uncertainty. Given that our 0.5° × 0.5° grid cell resolution can capture most of the population well in a given region, uncertainty associated with population was assumed to be negligible. We estimate the impacts of extra-regional emission reductions on mortality by using the response to extra-regional emission reduction (RERER) metric defined by TF HTAP (Galmarini et al., 2017):

$$\text{RERER}_i = \frac{R_{\text{global}} - R_{\text{region},i}}{R_{\text{global}}},$$

where for a given region $i$, $R_{\text{global}}$ is the change in mortality in the GLO 20% reduction simulation relative to the base simulation, and $R_{\text{region},i}$ is the change in mortality in response to the 20% emission reduction from that same region $i$. A RERER value near 1 indicates a strong relative influence of foreign emissions on mortality within a region, while a value near 0 indicates a weak foreign influence. We also estimate the total avoided extra-regional mortality from a source perspective as the sum of avoided deaths outside of each of the six source regions, and from a receptor perspective by summing $R_{\text{global}} - R_{\text{region},i}$ for all six regions.

3 Results

3.1 Response of O$_3$ and PM$_{2.5}$ concentrations to 20% regional and sectoral emission reductions

Previous TF HTAP studies reported area-averaged concentrations to quantify source–receptor relationships, averaging concentrations over a region (Doherty et al., 2013; Fiore et al., 2009; Fry et al., 2012; Huang et al., 2017; Stjern et al., 2016; Yu et al., 2013). Here, we present the population-weighted concentration over a region, which is more relevant for health. Among six receptor regions, the population-weighted multi-model mean O$_3$ concentrations range from 48.38 ± 8.05 ppb in EUR to 65.72 ± 10.08 ppb in SAS with a global average of 53.74 ± 8.03 ppb, while the annual population-weighted multi-model mean PM$_{2.5}$ concentrations range from 9.36 ± 2.62 µg m$^{-3}$ in NAM to 39.27 ± 13.50 µg m$^{-3}$ in EAS with a global average of 25.98 ± 5.05 µg m$^{-3}$ (Tables 1 and S5–S6 and Figs. S12–S13).

For 20% perturbation scenarios, in general the impact on the multi-model mean change in surface O$_3$ and PM$_{2.5}$ concentration is greater within the source region (i.e., domestic region) than outside of it (i.e., foreign region) (Figs. 1–2). This is also true for individual model results (Figs. S14–S16). Among six source regions, the emission reduction from SAS has the greatest impact on global population-weighted O$_3$ concentration (Tables 2 and S5), while that from EAS has the greatest impact on PM$_{2.5}$ (Tables 3 and S6). The source–receptor pairs with the greatest changes in O$_3$ and PM$_{2.5}$ concentration reflect the geographical proximity among regions and the magnitude of emissions (Tables 2–3) – e.g., EUR → MDE (0.34 ± 0.08 ppb), EUR → RBU (0.34 ppb ± 0.09), EAS → NAM (0.29 ± 0.14 ppb), EAS → RBU (0.27 ± 0.12 ppb), and NAM → EUR (0.26 ± 0.55 ppb) for O$_3$, and EUR → RBU (0.26 ± 0.19 µg m$^{-3}$), EUR → MDE (0.18 ± 0.08 µg m$^{-3}$), MDE → SAS (0.12 ± 0.06 µg m$^{-3}$), SAS → EAS (0.08 ± 0.08 µg m$^{-3}$), and EAS → SAS (0.08 ± 0.07 µg m$^{-3}$) for PM$_{2.5}$. Our ensemble shows ozone responses in the western US to emission reductions from EAS (Fig. 1c) similar to those modeled by Lin et al. (2012, 2017), who show that a model can capture the measured western US ozone increases due to rising Asian emissions.

For each receptor region, reducing foreign anthropogenic emissions by 20% (estimated by global minus within-region reductions) can decrease population-weighted O$_3$ concentrations by 29–74% of the change in O$_3$ concentration and 8–41% of the change in PM$_{2.5}$ concentration (Tables 2–3).
Table 1. Population-weighted multi-model mean O₃ (ppb) and PM₂.₅ concentration (µg m⁻³) for the 2010 baseline, for the 6-month O₃ season average of 1 h daily maximum O₃ and annual average PM₂.₅, shown with the standard deviation among models.

<table>
<thead>
<tr>
<th>Scenarios</th>
<th>Receptor regions</th>
</tr>
</thead>
<tbody>
<tr>
<td>O₃ (11 models)</td>
<td>NAM</td>
</tr>
<tr>
<td>PM₂.₅ (eight models)</td>
<td>9.36 ± 2.62</td>
</tr>
</tbody>
</table>

Table 2. Population-weighted multi-model mean change in O₃ (ppb) in receptor regions due to 20 % regional (NAM, EUR, SAS, MDE, and RBU), sectoral (PIN, TRN, and RES), and global (GLO) anthropogenic emission reductions, for the 6-month O₃ season average of 1 h daily maximum. The diagonal, showing the effect of each region on itself, is shown in italics. All numbers are rounded to the nearest hundredth, and are shown with standard deviations among models.

<table>
<thead>
<tr>
<th>Source regions/ sectors</th>
<th>Receptor region</th>
</tr>
</thead>
<tbody>
<tr>
<td>NAM</td>
<td>EUR</td>
</tr>
<tr>
<td>NAM</td>
<td>−1.88 ± 0.06</td>
</tr>
<tr>
<td>EUR</td>
<td>−0.08 ± 0.04</td>
</tr>
<tr>
<td>SAS</td>
<td>−0.05 ± 0.02</td>
</tr>
<tr>
<td>EAS</td>
<td>−0.29 ± 0.14</td>
</tr>
<tr>
<td>MDE</td>
<td>−0.04 ± 0.02</td>
</tr>
<tr>
<td>RBU</td>
<td>−0.05 ± 0.04</td>
</tr>
<tr>
<td>PIN</td>
<td>−1.13 ± 0.28</td>
</tr>
<tr>
<td>TRN</td>
<td>−1.26 ± 0.42</td>
</tr>
<tr>
<td>RES</td>
<td>−0.24 ± 0.09</td>
</tr>
<tr>
<td>GLO</td>
<td>−2.86 ± 0.77</td>
</tr>
</tbody>
</table>

In some cases, regional emission reductions cause small O₃ concentration increases within the source region or in foreign receptors, reflecting O₃ nonlinear responses (Fig. S14). For instance, C-IFS_v2 predicts O₃ concentration increases in EUR by 0.04 ppb from domestic emission reductions, which is in agreement with results from TF HTAP1 (Anenberg et al., 2009). Similarly, CAM-Chem shows more local O₃ increases, particularly in SAS, than other models (Fig. S14). The change in O₃ concentration in foreign receptors is broader than for PM₂.₅, reflecting that O₃ has a longer atmospheric lifetime than PM₂.₅.

For sectors, TRN emission reductions cause the greatest decrease in global population-weighted O₃ by 1.13 ± 0.19 ppb, while PIN emission reductions cause the greatest decrease in surface PM₂.₅ by 1.46 ± 0.56 µg m⁻³ globally (Tables 2–3 and Figs. 1–2). The 20 % emission reductions from individual sectors also have different effects in different regions. Of the three sectors, emission reductions from TRN have the greatest effect on population-weighted O₃ in NAM, EUR, SAS, MDE, and MDE (40–50 % of the global emission reduction) while PIN emission reductions dominate in EAS (57 %). Emission reductions from PIN have the greatest effect on population-weighted PM₂.₅ in NAM, EUR, EAS, MDE, and MDE (41–84 %) while RES emission reductions dominate in SAS (43 %). The response of PM₂.₅ concentration to sectoral emission reductions differs significantly across models, which reflects in part the PM₂.₅ species simulated by each model (Table S1 and Figs. S15–S17). For instance, we found that models that simulate PM₂.₅ nitrate (i.e., CHASER_t42 and GEOS-Chem Adjoint) predict a greater impact on PM₂.₅ concentration from TRN emission reduction than those without nitrate (i.e., GOCARTv5 and SPRINTARS) (Fig. S17).

3.2 Global mortality burden associated with anthropogenic air pollution

Table 4 shows the annual multi-model mean O₃- and PM₂.₅-related premature deaths in six regions and globally for the year 2010 baseline with 95 % CIs based on Monte Carlo sampling. Tables S7–S8 show estimates of premature deaths due to anthropogenic O₃ and PM₂.₅ from individual models. For the ensemble model mean, we estimate 290 000 (30 000, 600 000) premature O₃-related deaths globally using a 37.6 ppb counterfactual concentration, and 2.8 million (0.5 million, 4.6 million) PM₂.₅-related premature deaths using a uniform distribution of counterfactual concentration.
Figure 1. Global difference in multi-model mean O₃ concentrations (ppb) in 20% emission reduction scenarios relative to the baseline for the year 2010 in (a) North America (NAM), (b) Europe (EUR), (c) East Asia (EAS), (d) South Asia (SAS), (e) the Middle East (MDE), and (f) Russia–Belarus–Ukraine (RBU), the (g) power and industry (PIN), (h) transportation (TRN), and (i) residential (RES) sectors, and (j) globally (GLO), shown for the 6-month O₃ season average of 1 h daily maximum health-relevant metric.
Figure 2. Global difference in multi-model annual mean PM$_{2.5}$ concentrations ($\mu$g m$^{-3}$) in 20% emission reduction scenarios relative to the baseline for the year 2010 in (a) North America (NAM), (b) Europe (EUR), (c) East Asia (EAS), (d) South Asia (SAS), (e) the Middle East (MDE), and (f) Russia–Belarus–Ukraine (RBU), the (g) power and industry (PIN), (h) transportation (TRN), and (i) residential (RES) sectors, and (j) globally (GLO).
from 5.8 to 8.8 $\mu$g m$^{-3}$. Highly populated areas of India and EAS have the highest number of O$_3$- and PM$_{2.5}$-related deaths, and those regions together account for 82 and 66% of the global total O$_3$- and PM$_{2.5}$-related deaths. Compared with the GBD 2015 (Cohen et al., 2017), our global burden estimates are greater than the 254,000 (97,000, 422,000) premature deaths per year for O$_3$ from GBD, while there are less than 42 million (3.7 million, 4.8 million) premature deaths for PM$_{2.5}$, Lelieveld et al. (2015) estimate 142,000 (CI: 90,000, 208,000) O$_3$-related deaths and 3.2 million (1.5 million, 4.6 million) PM$_{2.5}$-related premature deaths for 2015. These differences can be explained mainly by exposure estimates. Here we used a multi-model ensemble, whereas Lelieveld et al. (2015) used a single model, and Cohen et al. (2017) used a single model for O$_3$ and a single model combined with surface and satellite observations for PM$_{2.5}$. In addition, Cohen et al. (2017) use RRs for particulate matter for IHD and stroke mortality that are modified from those used by Burnett et al. (2014) and applied age modification to the RRs, fitting the IER model for each age group separately. The updated IER with estimated higher relative risks, together with greater global pollution and baseline mortality rates in the low-income and middle-income countries in EAS and SAS, leads to the higher absolute numbers of attributable deaths and disability-adjusted life-years (DALYs) in GBD 2015 than estimated in GBD 2013 (Forouzanfar et al., 2016). Also, GBD 2015 includes lower child respiratory infection estimates whereas we do not. Our wider range of uncertainty for the global mortality reflects the uncertainty in baseline rates, RRs, and spread of air pollutant concentration across models whereas Cohen et al. (2017) consider national-level population-weighted mean concentrations and uncertainty of IER function predictions at each concentra-

tion, and Lelieveld et al. (2015) only account for the statistical uncertainty of the parameters used in the IER functions.

### 3.3 Effect of regional reductions on mortality

Reducing global anthropogenic emissions of air pollutant by 20% avoids 47,400 (11,300, 99,000) O$_3$-related deaths and 290,000 (67,100, 405,000) PM$_{2.5}$-related premature deaths (Tables 5–6 and S9–S10). Most avoided air-pollution-related deaths were found within or close to the source region (Figs. 3–6). Reducing anthropogenic emissions by 20% from NAM, EUR, SAS, EAS, MDE, and RBU can avoid 54, 54, 95, 85, 21, and 22% of the global change in O$_3$-related deaths within the source region (the number of avoided deaths within the source region is divided by the number of avoided deaths globally), and 93, 81, 93, 94, 32, and 82% of the global change in PM$_{2.5}$-related deaths, respectively (Tables 5–6). Whereas the most O$_3$-related premature deaths can be avoided by reducing SAS emissions (20,000 (3,600, 42,200) deaths per year), reducing EAS emissions avoids more O$_3$-related premature deaths (1700 (−1300, 3500) deaths per year) outside of the source region than for any other region (500 (180, 870) deaths per year to 1300 (−1200, 4400) deaths per year (Table 5)). Similarly, while reducing EAS emissions avoids the most PM$_{2.5}$-related premature deaths (96,600 (3500, 136,000) deaths per year), reducing EUR emissions avoids more PM$_{2.5}$-related premature deaths (7400 (930, 9500) deaths per year) outside of the source region than for any other region (1400 (−320, 2300) deaths per year to 5500 (3000, 7800) deaths per year) (Table 6). While emission reductions from one region generally lead to more avoided deaths within the source region than outside of it, 20% anthropogenic emission reductions from MDE (i.e., 79 and 68% of global avoided deaths outside the source region for O$_3$ and PM$_{2.5}$, respectively) and RBU (78% for

<table>
<thead>
<tr>
<th>Source regions/sectors</th>
<th>NAM</th>
<th>EUR</th>
<th>SAS</th>
<th>EAS</th>
<th>MDE</th>
<th>RBU</th>
<th>World</th>
</tr>
</thead>
<tbody>
<tr>
<td>NAM</td>
<td>−1.32 ± 0.66</td>
<td>−0.03 ± 0.02</td>
<td>0.00 ± 0.01</td>
<td>−0.02 ± 0.02</td>
<td>−0.01 ± 0.01</td>
<td>−0.01 ± 0.01</td>
<td>−0.08 ± 0.04</td>
</tr>
<tr>
<td>EUR</td>
<td>−0.01 ± 0.00</td>
<td>−1.17 ± 0.87</td>
<td>−0.01 ± 0.01</td>
<td>−0.02 ± 0.01</td>
<td>−0.18 ± 0.08</td>
<td>−0.26 ± 0.19</td>
<td>−0.13 ± 0.09</td>
</tr>
<tr>
<td>SAS</td>
<td>&lt; −0.01</td>
<td>&lt; −0.01</td>
<td>&lt; −4.86 ± 2.17</td>
<td>−0.08 ± 0.08</td>
<td>−0.03 ± 0.02</td>
<td>&lt; −0.01</td>
<td>−1.16 ± 0.51</td>
</tr>
<tr>
<td>EAS</td>
<td>−0.03 ± 0.01</td>
<td>−0.02 ± 0.01</td>
<td>−0.08 ± 0.07</td>
<td>−6.19 ± 2.08</td>
<td>&lt; −0.01</td>
<td>−0.04 ± 0.02</td>
<td>−1.45 ± 0.71</td>
</tr>
<tr>
<td>MDE</td>
<td>&lt; −0.01</td>
<td>&lt; −0.01</td>
<td>&lt; −0.12 ± 0.06</td>
<td>−0.01 ± 0.02</td>
<td>−0.91 ± 0.38</td>
<td>−0.05 ± 0.03</td>
<td>−0.08 ± 0.03</td>
</tr>
<tr>
<td>RBU</td>
<td>&lt; −0.01</td>
<td>&lt; −0.01</td>
<td>&lt; −0.07 ± 0.05</td>
<td>−0.01 ± 0.02</td>
<td>−0.04 ± 0.02</td>
<td>−0.03 ± 0.02</td>
<td>−0.78 ± 0.50</td>
</tr>
<tr>
<td>PIN</td>
<td>−0.61 ± 0.18</td>
<td>−0.57 ± 0.26</td>
<td>−1.73 ± 0.71</td>
<td>−2.75 ± 0.99</td>
<td>−0.92 ± 0.14</td>
<td>−0.58 ± 0.19</td>
<td>−1.46 ± 0.56</td>
</tr>
<tr>
<td>TRN</td>
<td>−0.27 ± 0.20</td>
<td>−0.38 ± 0.41</td>
<td>−0.82 ± 0.88</td>
<td>−0.54 ± 0.43</td>
<td>−0.09 ± 0.06</td>
<td>−0.15 ± 0.16</td>
<td>−0.40 ± 0.37</td>
</tr>
<tr>
<td>RES</td>
<td>−0.20 ± 0.05</td>
<td>−0.27 ± 0.12</td>
<td>−1.93 ± 0.40</td>
<td>−1.70 ± 0.28</td>
<td>−0.08 ± 0.02</td>
<td>−0.20 ± 0.05</td>
<td>−1.17 ± 0.31</td>
</tr>
<tr>
<td>GLO</td>
<td>−1.47 ± 0.72</td>
<td>−1.52 ± 1.04</td>
<td>−5.40 ± 2.31</td>
<td>−6.76 ± 3.29</td>
<td>−1.55 ± 0.75</td>
<td>−1.19 ± 0.73</td>
<td>−3.49 ± 1.51</td>
</tr>
</tbody>
</table>
Table 4. Annual multi-model empirical mean O$_3$- and PM$_{2.5}$-related premature deaths with 95% CI from Monte Carlo simulations in parentheses (including uncertainty in baseline mortality rates, RRs, and air pollutant concentration across models) in the year 2010 baseline. All numbers are rounded to three significant figures or the nearest 100 deaths. Empirical mean is the mean of 1000 Monte Carlo simulations.

<table>
<thead>
<tr>
<th>Receptor region</th>
<th>NAM</th>
<th>EUR</th>
<th>SAS</th>
<th>EAS</th>
<th>MDE</th>
<th>RBU</th>
<th>World</th>
</tr>
</thead>
<tbody>
<tr>
<td>O$_3$ (11 models)</td>
<td>(900–30 000)</td>
<td>(600–28 000)</td>
<td>(23 000–277 000)</td>
<td>(3900–213 000)</td>
<td>(300–7000)</td>
<td>(100–6600)</td>
<td>(30 000–596 000)</td>
</tr>
<tr>
<td>PM$_{2.5}$ (eight models)</td>
<td>(1500–158 000)</td>
<td>(2700–463 000)</td>
<td>(328 000–1 110 000)</td>
<td>(159 000–1 720 000)</td>
<td>(600–133 000)</td>
<td>(2700–358 000)</td>
<td>(514 000–4 640 000)</td>
</tr>
</tbody>
</table>

O$_3$) can avoid more premature deaths outside of the source region than within (Tables 5–6). This result for RBU is in agreement with West et al. (2009b). However, the results for NAM and EUR do not agree with previous studies that found that emission reductions in these regions cause more O$_3$-related avoided premature deaths outside of the source region than within (Anenberg et al., 2009; Duncan et al., 2008; West et al., 2009b). For PM$_{2.5}$, our results are comparable with Anenberg et al. (2014) and Crippa et al. (2017), who found that for most regions, PM$_{2.5}$-related avoided premature deaths are higher within the source region than outside. The above difference in results with TF HTAP1 may be in part because of the definition of regions. Whereas the TF HTAP2 regions are defined by geopolitical boundaries, the TF HTAP1 regions are defined by square domains which are larger and include more ocean areas (Anenberg et al., 2009). In addition, updated atmospheric models and emission inputs, as well as different atmospheric dynamics in the single years chosen in TF HTAP1 vs. TF HTAP2 may contribute to the differences.

Using individual models, different conclusions may result for the relative importance of interregional transport. For example, for O$_3$, eight models predict that NAM emission reductions cause more O$_3$-related premature deaths within NAM (i.e., CAM-Chem, CHASER_T42, CHASER_T106, C-IFS, GEOS-Chem Adjoint, GEOS-Chem, GFDL_AM3, and HadGEM2-ES), whereas two models predict more deaths outside NAM (i.e., EMEPv48 and OsloCTM3.v2). Five models suggest that EUR emission reductions cause more O$_3$-related premature deaths within EUR (i.e., CAM-Chem, CHASER_T42, CHASER_T106, GFDL_AM3, and HadGEM2-ES), whereas four show more deaths outside (i.e., C-IFS, GEOS-Chem Adjoint, EMEPv48, and OsloCTM3.v2). Each individual model shows that emission reductions from SAS and EAS avoid more O$_3$-related premature deaths within than outside, and that those from MDE and RBU avoid more O$_3$-related premature deaths outside than within (Fig. S18). For PM$_{2.5}$, each individual model shows that emission reductions from NAM, EUR, SAS, EAS, and RBU avoid more PM$_{2.5}$-related premature deaths within than outside, while for emission reductions from MDE, three models (EMEPv48, GEOS-Chem Adjoint, and SPRINT-ARS) show more PM$_{2.5}$-related premature deaths within, while three (CHASER_T42, GEOS5, and GOCART) show more PM$_{2.5}$-related premature deaths outside (Fig. S19). The variation in health effect reflects the differences in processing of natural emissions, atmospheric physical and chemical mechanisms, numerics, etc., across models.

For each receptor region, reducing domestic anthropogenic emissions by 20% contributes about 66, 39, 84, 72, 45, and 25% of the total O$_3$-related avoided premature mortality (from the global reduction) and 90, 78, 87, 58, and 66% of the total PM$_{2.5}$-related avoided premature mortality (from the global reduction) in NAM, EUR, SAS, EAS, MDE, and RBU, respectively (Tables 5–6). Therefore, reducing emissions from foreign regions avoids more O$_3$ premature deaths in EUR (foreign emissions account for 61% of total avoided deaths from the global reduction), MDE (55%), and RBU (75%) than reducing domestic emissions (Tables 5–6), in agreement with the results for EUR from Anenberg et al. (2009). Whereas EAS has the greatest number of avoided O$_3$-related premature deaths due to foreign emission reduction (3800 (3600, 3900) deaths per year), RBU has the greatest fraction of O$_3$ mortality from foreign emission reductions (75%) (Table 5). Similarly, for PM$_{2.5}$, while EAS has the greatest number of avoided PM$_{2.5}$-related premature deaths due to foreign emission reductions (13 600 (3500, 18 800) deaths per year), MDE has the greatest fraction of PM$_{2.5}$ mortality from foreign emission reduction (42%) (Table 6).

Overall, adding results from all six regional reductions, interregional transport of air pollution from extra-regional contributions is estimated to lead to more avoided deaths through changes in PM$_{2.5}$ (25 100 (8200, 35 800) deaths per year) than in O$_3$ (6000 (−3400, 15 500) deaths per year), consistent with Anenberg et al. (2009, 2014). This result is due to the greater influence of PM$_{2.5}$ on mortality, despite the shorter atmospheric lifetime of PM$_{2.5}$ relative to O$_3$.

The contributions of different factors to the overall uncertainties in mortality are shown in Tables S11–S12, considering uncertainties due to the spread of air pollutant concentrations across models, RRs, and baseline mortality rates, expressed as coefficients of variation. For both O$_3$ and PM$_{2.5}$ mortality, the spread of model results generally contributes most to the overall uncertainty, followed by uncertainty in RRs and in baseline mortality rates, for most source–receptor pairs. The spread of model results is generally wider for
Figure 3. Annual avoided O$_3$-related premature deaths in 2010 per 1000 km$^2$ due to 20% emission reduction scenarios relative to the base case in (a) North America (NAM), (b) Europe (EUR), (c) East Asia (EAS), (d) South Asia (SAS), (e) the Middle East (MDE), and (f) Russia–Belarus–Ukraine (RBU), the (g) power and industry (PIN), (h) transportation (TRN), and (i) residential (RES) sectors, and (j) globally (GLO).
together avoids the highest number of O$_3$ reducing anthropogenic emissions by 20 % in three sectors (e.g., NAM, EUR, India, China) (Figs. 3–6). For instance, air-pollution-related premature deaths in highly populated areas (e.g., NAM, EUR, SAS, MDE, and RBU), sectoral (PIN, TRN, and RES), and global (GLO) anthropogenic emission reductions in each region and worldwide. The diagonal, showing the effect of each region on itself, is shown in italics. For regional reductions, we also use the RERER (Eq. 4) as the percent of total avoided deaths in each receptor region that result from foreign emission reductions, as well as the percent of global avoided deaths from emission reductions in each source region. All numbers are rounded to three significant figures or the nearest 10 deaths.

### Table 5

<table>
<thead>
<tr>
<th>Source regions/sectors</th>
<th>NAM</th>
<th>EUR</th>
<th>SAS</th>
<th>EAS</th>
<th>MDE</th>
<th>RBU</th>
<th>Worldwide</th>
</tr>
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<tbody>
<tr>
<td>NAM</td>
<td>1500</td>
<td>330</td>
<td>170</td>
<td>500</td>
<td>30</td>
<td>70</td>
<td>2,800</td>
</tr>
<tr>
<td></td>
<td>(−170–4000)</td>
<td>(10–780)</td>
<td>(−250–690)</td>
<td>(−910–2200)</td>
<td>(0–80)</td>
<td>(0–170)</td>
<td>(−1300–8400)</td>
</tr>
<tr>
<td>EUR</td>
<td>60</td>
<td>930</td>
<td>−80</td>
<td>490</td>
<td>50</td>
<td>110</td>
<td>1,700</td>
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<tr>
<td>SAS</td>
<td>40</td>
<td>50</td>
<td>19,000</td>
<td>420</td>
<td>20</td>
<td>10</td>
<td>20,000</td>
</tr>
<tr>
<td></td>
<td>(−40–130)</td>
<td>(−30–160)</td>
<td>(4000–42,000)</td>
<td>(−340–1400)</td>
<td>(0–40)</td>
<td>(−10–40)</td>
<td>(3600–42,200)</td>
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<tr>
<td>EAS</td>
<td>230</td>
<td>310</td>
<td>450</td>
<td>9700</td>
<td>30</td>
<td>80</td>
<td>11,400</td>
</tr>
<tr>
<td>MDE</td>
<td>30</td>
<td>60</td>
<td>310</td>
<td>160</td>
<td>180</td>
<td>30</td>
<td>870</td>
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<td></td>
<td>(−30–120)</td>
<td>(−50–190)</td>
<td>(−90–910)</td>
<td>(−120–520)</td>
<td>(−10–480)</td>
<td>(0–70)</td>
<td>(−330–2600)</td>
</tr>
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<td>RBU</td>
<td>40</td>
<td>150</td>
<td>−200</td>
<td>420</td>
<td>20</td>
<td>140</td>
<td>640</td>
</tr>
<tr>
<td>PIN</td>
<td>900</td>
<td>850</td>
<td>7400</td>
<td>7800</td>
<td>140</td>
<td>210</td>
<td>19,300</td>
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<tr>
<td>TRN</td>
<td>1000</td>
<td>970</td>
<td>10,600</td>
<td>3500</td>
<td>210</td>
<td>200</td>
<td>18,800</td>
</tr>
<tr>
<td>RES</td>
<td>200</td>
<td>250</td>
<td>6000</td>
<td>3000</td>
<td>30</td>
<td>60</td>
<td>10,400</td>
</tr>
<tr>
<td></td>
<td>(−20–510)</td>
<td>(40–550)</td>
<td>(1600–12,200)</td>
<td>(670–6300)</td>
<td>(0–80)</td>
<td>(10–120)</td>
<td>(2700–21,100)</td>
</tr>
<tr>
<td>GLO</td>
<td>2300</td>
<td>2400</td>
<td>22,600</td>
<td>13,500</td>
<td>400</td>
<td>550</td>
<td>47,400</td>
</tr>
<tr>
<td>RERER</td>
<td>34 %</td>
<td>61 %</td>
<td>16 %</td>
<td>28 %</td>
<td>55 %</td>
<td>75 %</td>
<td>–</td>
</tr>
</tbody>
</table>

PM$_{2.5}$ (14 to 3974 % among source–receptor pairs) than for O$_3$ (13 to 1065 %). The uncertainty in RRs for O$_3$ mortality has a constant value (33 to 34 %) due to the fixed uncertainty range of RRs from Jerrett et al. (2009), whereas PM$_{2.5}$ mortality leads to a wider range of uncertainty (1 to 247 %) in RRs because the uncertainty differs at different PM$_{2.5}$ concentrations (Burnett et al., 2014). Low uncertainty in baseline mortality rate was found for most source–receptor pairs (< 20 %) except for the response of PM$_{2.5}$ mortality in SAS to 20 % reduction from RBU (66 %).

### 3.4 Effect of sectoral reductions on mortality

Reducing global anthropogenic emissions by 20 % in three sectors (i.e., PIN, TRN, and RES) together avoids 48,500 (7100, 108,000) O$_3$-related premature deaths and 243,000 (66,800, 357,000) PM$_{2.5}$-related premature deaths globally (Tables 5–6), with the greatest number of avoided air-pollution-related premature deaths in highly populated areas (e.g., NAM, EUR, India, China) (Figs. 3–6). For instance, reducing anthropogenic emissions by 20 % in three sectors together avoids the highest number of O$_3$-related deaths in SAS (24,000 (6000, 49,600) deaths per year) and PM$_{2.5}$-related deaths in EAS (83,400 (29,400, 135,000) deaths per year). We compare our estimates of O$_3$- and PM$_{2.5}$-related premature deaths attributable to PIN, TRN, and RES emissions with previous studies by multiplying our results for 20 % emission reductions by 5 and by combining their sectors to nearly match each of the three sectors in this study (Table 7). Compared with Silva et al. (2016a), our estimate of O$_3$- and PM$_{2.5}$-related premature deaths attributable to PIN and TRN are very comparable, but that attributable to RES is lower here. In comparison with Lelieveld et al. (2015), we estimate greater O$_3$- and PM$_{2.5}$-related premature deaths attributable to PIN and TRN, but fewer for RES.

Like Silva et al. (2016a) and Lelieveld et al. (2015), different locations show relatively different mortality responses to changes in sectoral emissions. Whereas PIN emission reductions cause the greatest number of avoided O$_3$-related premature deaths globally (19,300 (1400, 45,000) deaths per year), TRN emission reductions cause the greatest fraction of avoided deaths in most of the six regions (26–53 % of the global emission reduction), except for EAS (58 %) and RBU (38 %), where the effect of reducing PIN emissions dominates. In comparison with other studies (Table 7), our con-

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Atmos. Chem. Phys., 18, 10497–10520, 2018 www.atmos-chem-phys.net/18/10497/2018

For PM\textsubscript{2.5} deaths in most regions agrees well with Silva et al. (2016a). and TRN emissions cause the greatest fraction of avoided emissions from this study (for 20\% reductions) are multiplied by 5. For Silva et al. (2016), we combine results for “energy” and “industry” to represent RES. For Lelieveld et al. (2015), we use the RERER (Eq. 4) as the percent of total avoided deaths in each receptor region that result from foreign emission reductions, as well as the percent of global avoided deaths from emission reductions in each source region. All numbers are rounded to three significant figures or the nearest 10 deaths.

Table 6. Annual avoided multi-model empirical mean PM\textsubscript{2.5}-related premature deaths (IHD + STROKE + COPD + LC) with 95\% CI from Monte Carlo simulations in parentheses due to 20\% regional (NAM, EUR, SAS, MDE, and RBU), sectoral (PIN, TRN, and RES), and global (GLO) anthropogenic emission reductions in each region and worldwide. The diagonal, showing the effect of each region on itself, is shown in italics. For reductions, we also use the RERER (Eq. 4) as the percent of total avoided deaths in each receptor region that result from foreign emission reductions, as well as the percent of global avoided deaths from emission reductions in each source region. All numbers are rounded to three significant figures or the nearest 10 deaths.

<table>
<thead>
<tr>
<th>Source regions</th>
<th>Receptor region</th>
<th>Impact on foreign receptor regions</th>
</tr>
</thead>
<tbody>
<tr>
<td>NAM (60–29 800)</td>
<td>EUR (80–1100)</td>
<td>19 400 (310–30 600)</td>
</tr>
<tr>
<td>EUR (20–110)</td>
<td>SAS (0–200)</td>
<td>39 400 (5 500–63 400)</td>
</tr>
<tr>
<td>SAS (0–10)</td>
<td>EAS (30–3700)</td>
<td>51 300 (32 300–73 300)</td>
</tr>
<tr>
<td>EAS (40–510)</td>
<td>MDE (0–200)</td>
<td>96 600 (35 000–136 000)</td>
</tr>
<tr>
<td>MDE (0–60)</td>
<td>RBU (90–850)</td>
<td>5000 (1900–11 100)</td>
</tr>
<tr>
<td>RBU (10–60)</td>
<td>PIN (940–13 000)</td>
<td>21 500 (900–31 000)</td>
</tr>
<tr>
<td>PIN (940–13 000)</td>
<td>TRN (0–1100)</td>
<td>128 000 (41 600–179 000)</td>
</tr>
<tr>
<td>TRN (0–3700)</td>
<td>RES (10–60)</td>
<td>83 400 (41 600–120 000)</td>
</tr>
<tr>
<td>RES (710–31 300)</td>
<td>GLO (4900–68 100)</td>
<td>290 000 (67 100–405 000)</td>
</tr>
<tr>
<td>RERER</td>
<td>22%</td>
<td>13%</td>
</tr>
</tbody>
</table>

Table 7. Comparison of O\textsubscript{3}- and PM\textsubscript{2.5}-related premature deaths attributable to PIN, TRN, and RES emissions with previous studies. Results from this study (for 20\% reductions) are multiplied by 5. For Silva et al. (2016), we combine results for “energy” and “industry” to represent PIN, and use “land transportation” to represent TRN and “residential & commercial” to represent RES. For Lelieveld et al. (2015), we combine the “power generation” and “industry” sectors to represent PIN, and use “land traffic” to represent TRN and “residential” energy to represent RES.

<table>
<thead>
<tr>
<th>Emission source sector</th>
<th>This study</th>
<th>Silva et al. (2016)</th>
<th>Lelieveld et al. (2015)</th>
</tr>
</thead>
<tbody>
<tr>
<td>PIN</td>
<td>O\textsubscript{3} : 96 500 (7000, 225 000)</td>
<td>O\textsubscript{3} : 111 000 (23 200, 240 000)</td>
<td>O\textsubscript{3} + PM\textsubscript{2.5} 692 000</td>
</tr>
<tr>
<td></td>
<td>PM\textsubscript{2.5} : 640 000 (208 000, 895 000)</td>
<td>PM\textsubscript{2.5} : 613 000 (422 000, 816 000)</td>
<td></td>
</tr>
<tr>
<td>TRN</td>
<td>O\textsubscript{3} : 94 000 (15 000, 208 000)</td>
<td>O\textsubscript{3} : 80 900 (17 400, 180 000)</td>
<td>O\textsubscript{3} + PM\textsubscript{2.5} 165 000</td>
</tr>
<tr>
<td></td>
<td>PM\textsubscript{2.5} : 160 000 (82 500, 292 000)</td>
<td>PM\textsubscript{2.5} : 212 000 (114 000, 292 000)</td>
<td></td>
</tr>
<tr>
<td>RES</td>
<td>O\textsubscript{3} : 52 000 (13 500, 106 000)</td>
<td>O\textsubscript{3} : 53 700 (12 300, 116 000)</td>
<td>O\textsubscript{3} + PM\textsubscript{2.5} 1 020 000</td>
</tr>
<tr>
<td></td>
<td>PM\textsubscript{2.5} : 417 000 (209 000, 600 000)</td>
<td>PM\textsubscript{2.5} : 675 000 (428 000, 899 000)</td>
<td></td>
</tr>
</tbody>
</table>

Conclusion that PIN emissions cause the most O\textsubscript{3}-related deaths and TRN emissions cause the greatest fraction of avoided deaths in most regions agrees well with Silva et al. (2016a). For PM\textsubscript{2.5}, reducing PIN emissions avoids the most PM\textsubscript{2.5}-related premature deaths globally (128 000 (41 600, 179 000) deaths per year) and in most regions (38–78\% of the global emission reduction), except for SAS (45\%), where the RES emissions dominate. Although these findings differ from those of Lelieveld et al. (2015) and Silva et al. (2016), who find that RES emissions have the greatest impact on PM\textsubscript{2.5} mortality globally and in most regions, all studies agree that PIN emissions have the greatest impact in NAM. Our result is
Figure 4. Annual avoided O$_3$-related premature deaths in 2010 per million people due to 20% emission reduction scenarios relative to the base case in (a) North America (NAM), (b) Europe (EUR), (c) East Asia (EAS), (d) South Asia (SAS), (e) the Middle East (MDE), and (f) Russia–Belarus–Ukraine (RBU), the (g) power and industry (PIN), (h) transportation (TRN), and (i) residential (RES) sectors, and (j) globally (GLO).
Figure 5. Annual avoided PM$_{2.5}$-related premature deaths in 2010 per 1000 km$^2$ due to 20% emission reduction scenarios relative to the base case in (a) North America (NAM), (b) Europe (EUR), (c) East Asia (EAS), (d) South Asia (SAS), (e) the Middle East (MDE), and (f) Russia–Belarus–Ukraine (RBU), the (g) power and industry (PIN), (h) transportation (TRN), and (i) residential (RES) sectors, and (j) globally (GLO).
Figure 6. Annual avoided PM$_{2.5}$-related premature deaths in 2010 per million people due to 20% emission reduction scenarios relative to the base case in (a) North America (NAM), (b) Europe (EUR), (c) East Asia (EAS), (d) South Asia (SAS), (e) the Middle East (MDE), and (f) Russia–Belarus–Ukraine (RBU), the (g) power and industry (PIN), (h) transportation (TRN), and (i) residential (RES) sectors, and (j) globally (GLO).
also comparable with Crippa et al. (2017), who find that PIN emissions have the greatest health impact in most countries. Although comparable emission inventories are used (i.e., Lelieveld et al. (2015) and this study use EDGAR emissions while Silva et al. (2016) use RCP8.5 emissions), our lower mortality estimate for RES emissions may be explained by our 20 % reductions relative to the zero-out method and the different years simulated.

Considering results from individual models, we found that mortality from TRN emission reductions shows greater relative uncertainty than from PIN or RES (Tables S5–S10), reflecting a greater spread of results across models. Regional impacts from individual models also differ from the ensemble mean result – e.g., for O3, GEOS-Chem Adjoint and OsloCTM3.v2 show that reducing PIN emissions causes the greatest fraction of avoided O3-related deaths in EUR, while GEOS-Chem Adjoint, HadGM2-ES, and OsloCTM3.v2 show that TRN emissions have the greatest fraction of avoided O3-related deaths in RBU (Fig. S20). For PM2.5, CHASER_t42 and GEOS-Chem Adjoint show that reducing PIN emissions causes the greatest fraction of avoided PM2.5-related deaths in SAS (Fig. S21).

4 Discussion

We aggregate the avoided deaths attributable to 20 % reductions from four corresponding source regions (i.e., NAM, EUR, SAS, and EAS) and compare with the findings from TF HTAP1. We estimate that these regional emission reductions are associated with 36 000 (−1500, 90 300) avoided deaths globally through the change in O3 and 207 000 (41 500, 304 000) avoided deaths through the change in PM2.5, more than those estimated by Anenberg et al. (2009, 2014) – 21 800 (10 600, 33 400) deaths for O3 and 192 000 (146 000, 230 000) deaths for PM2.5. This discrepancy might be attributed to different health impact functions, emission datasets, region definitions, updated population, or baseline mortality rates. In particular, for O3 respiratory mortality, we use a log-linear model for chronic mortality (Jerrrett et al., 2009), instead of the short-term O3 mortality estimate based on a daily time series study (Bell et al., 2004) used by Anenberg et al. (2009). For PM2.5 mortality, Anenberg et al. (2014) only included the simulated changes in BC, particulate organic matter (POM = primary organic aerosol + secondary organic aerosol), and sulfate for PM2.5 concentration, while we use the total model-reported PM2.5 concentration, which includes more species for some models. We also apply the IER model (Burnett et al., 2014) for PM2.5, as opposed to the log-linear model of Krewski et al. (2009) used by Anenberg et al. (2014).

For regional reductions, our multi-model average results suggest that NAM and EUR emissions cause more deaths inside of those regions than outside, which disagrees with previous studies (Anenberg et al., 2009; Duncan et al., 2008; West et al., 2009b), whereas similar regional impacts are found for EAS and SAS. Also, total avoided deaths through interregional air pollution transport are estimated as 6000 (−3400, 15 500) deaths per year for O3 and 25 100 (8200, 35 800) deaths per year for PM2.5 in this study, in contrast with 7300 (3600, 11 200) deaths per year for O3 and 11 500 (8800, 14 200) deaths per year for PM2.5 in Anenberg et al. (2009, 2014). These differences likely result from different concentration response functions and the use of six regions here vs. four by Anenberg et al. (2009, 2014). In addition, updated atmospheric models and emission inputs, as well as different atmospheric dynamics in the single years chosen in TF HTAP1 vs. TF HTAP2 may contribute to the differences. In addition, updated atmospheric models and emission inputs, as well as different atmospheric dynamics in the single years chosen in HTAP vs. HTAP2 may contribute to the differences. Overall, whereas O3 accounts for a higher percentage of the total deaths in foreign regions than PM2.5, PM2.5 leads to more deaths in general, which agrees well with the results of Anenberg et al. (2009, 2014).

Using regional models in AQMEII3, driven by a single global model (C-IFS_v2), Im et al. (2018) estimated that 20 % domestic emission reductions would avoid 54 000 and 27 500 premature deaths (for O3 and PM2.5 combined) in Europe and the US, respectively, as opposed to ~1000 and 2000 premature deaths due to foreign emission reductions. These results are comparable to our estimates that 32 900 and 19 500 premature deaths result from 20 % domestic emission reductions in Europe and the US, while 670 and 570 premature deaths result from foreign emission reductions. Although our defined US region is slightly bigger than Im et al. (2018), the majority of US emission sources and population are located within the region defined by Im et al. (2018). This comparison shows that regional and global models show similar impacts on mortality from air pollution transport.

Differences in our estimates of premature mortality attributable to air pollution from three emission sectors (multiplied by 5) may be explained by methodological differences relative to previous studies (Silva et al., 2016; Lelieveld et al., 2015), including our use of 20 % emission reductions versus the zero-out method in those studies, different emission inventories, a multi-model ensemble versus single models, and differences in baseline mortality rates, population, and concentration response functions. Our finding that TRN emissions contribute the most avoided deaths for O3 in most regions agrees well with the result by Silva et al. (2016a), but differs for PM2.5 mortality for which we find that PIN emissions cause the most deaths, while both Silva et al. (2016a) and Lelieveld et al. (2015) find that RES emissions are responsible for the most deaths. This discrepancy may be explained by different PM2.5 species included in individual models, as we showed that changes in PM2.5 concentration to TRN emission differ across models.

By using an ensemble of multi-model results here, we highlight the relative importance of different source–receptor
pairs for mortality in a way that is more robust than using a single model, particularly since some individual models yielded different conclusions than the ensemble mean. The air pollutant concentration changes reported by the HTAP2 models may be different among models; it may result from a variety of processes, e.g., atmospheric physical and chemical mechanisms, processing of natural emissions, and transport time step (Table S1), but not anthropogenic emissions since those were nearly identical among models. In addition, the coarse model resolution used by global models may underestimate health effects by misaligning peak concentration and population, particularly in urban areas and for PM$_{2.5}$ (Punger and West, 2013), but it is not known how model resolution would affect the relative contributions of extra-regional and intraregional health benefits. Future research should explore the possible bias from using coarse global models for extra-regional and intraregional mortality estimates in metropolitan regions by comparing with finer-resolution chemical transport models.

Another uncertainty in this paper (and other global studies) lies in applying the same RR$\text{s}$ worldwide because of lack of long-term records of the chronic influences of ambient air pollution on mortality outside of NAM and EUR. We consider only the population of adults $\geq$ 25 years old, ignoring possible mortality effects on the younger population, and consequently we may underestimate premature mortality overall. Likewise, the effects of air pollution on several morbidity end points are omitted. We assume that all PM$_{2.5}$ is equally toxic, for lack of clear evidence for greater toxicity of some species. Intergovernmental transport can also change the toxicity of PM$_{2.5}$ by changing the size distribution or chemical composition, where transport likely causes particles to become more oxidized (West et al., 2016). Future research on PM$_{2.5}$-related mortality should include estimating health effects for different PM$_{2.5}$ chemical components.

5 Conclusions

We estimate O$_3$- and PM$_{2.5}$-related premature mortality from simulations with 14 global chemistry transport models participating in the TF HTAP2 multi-model exercise for the year 2010. An estimate of 290,000 (30,000, 600,000) global premature O$_3$-related deaths and 2.8 million (0.5 million, 4.6 million) global PM$_{2.5}$-related premature deaths is obtained from the ensemble for the year 2010 in the baseline case. We focus on model experiments simulating 20% regional air pollutant emission reductions (excluding methane) in six regions, three sectors, and one global domain. For regional scenarios, six source emission reductions altogether can cause 84% of the global avoided O$_3$-related premature deaths within the source region, ranging from 21 to 95% among six regions, and 16% (5 to 79%) outside of the source region. For PM$_{2.5}$, 89% of global avoided PM$_{2.5}$-related premature deaths are within the source region, ranging from 32 to 94% among six regions, and 11% (6 to 68%) outside of the source region. While most avoided mortality generally occurs within the source region, we find that emission reductions from RBU (only for O$_3$) and MDE (for both O$_3$ and PM$_{2.5}$) can avoid more premature deaths outside of these regions than within. Considering the effects of foreign emissions on receptor regions, 20% foreign emission reductions lead to more avoided O$_3$-related premature deaths in EUR, MDE, and RBU than domestic reductions. Reductions from all six regions in the transport of air pollution among regions are estimated to lead to more avoided deaths through changes in PM$_{2.5}$ (25,100 (8,200, 35,800) deaths per year) than for O$_3$ (6,000 (−3,400, 15,500) deaths per year). For NAM and EUR, our estimates of avoided mortality from regional and extra-regional emission reductions are comparable to those estimated by regional models in AQMEII3 (Im et al., 2018) for these same emission reduction experiments. Overall, the spread of modeled air pollutant concentrations contributes most to the uncertainty in mortality estimates, highlighting that using a single model may lead to erroneous conclusions and may underestimate uncertainty in mortality estimates.

For sectoral emission reductions, reducing anthropogenic emissions by 20% in three sectors together avoids 48,500 (7100, 108,000) O$_3$-related premature deaths and 243,000 (66,800, 357,000) PM$_{2.5}$-related premature deaths globally. Of the three sectors, TRN had the greatest fraction (26–53%) of O$_3$-related premature deaths globally and in most regions, except for EAS (58%) and RBU (38%) where PIN emissions dominate. For PM$_{2.5}$ mortality, PIN emissions cause the most deaths in most regions (38–78%), except for SAS (45%) where RES emissions dominate.

In this study, we have gone beyond previous TF HTAP1 studies that quantified premature mortality from interregional air pollution transport by using more source regions, analyzing source emission sectors, and using updated atmospheric models and health impact functions. The estimate of air transport premature mortality could vary due to differences in exposure estimate (single model vs. ensemble model), health impact function, regional definitions, and grid resolutions. These discrepancies highlight uncertainty estimated by different methods in previous studies. Despite uncertainties, our results suggest that reducing pollution transported over a long distance would be beneficial for health, with impacts from all foreign emission reductions combined that may be comparable to or even exceed the impacts of emission reductions within a region. Additionally, actions to reduce emissions should target specific sectors within world regions, as different sectors dominate the health effects in different regions. This work highlights the importance of long-range air pollution transport and suggests that estimates of the health benefits of emission reductions on local, national, or continental scales may underestimate the overall health benefits globally, when interregional transport is accounted for. International cooperation to reduce air pollution transported over long distances may therefore be desirable.
Author contributions. CKL collected all model output used in this supplement.

Information about the Supplement

A detailed description of the models participating in the ensemble, a map of six priority regions used in this analysis, and additional results can be found in the Supplement.

The Supplement related to this article is available online at https://doi.org/10.5194/acp-18-10497-2018-supplement.

Author contributions. CKL collected all model output used in this study, performed health impact calculations, and wrote the paper. JJW conceived of the study and supervised CKL. RAS developed the health assessment code and preprocessed the required health input data. CAM-Chem model experiments were prepared by LE. CHASER_T42 and CHASER_T106 model experiments were prepared by KS and TS. C-IFS model experiments were prepared by JF. EMEPv48 model experiments were prepared by JEJ. GEOS5 model experiments were prepared by HB, MC, and XP. GEOSCHEMADJOINT model experiments were prepared by DH and YD. GEOS-Chem model experiments were prepared by RJP. GFDL_AM3 model experiments were prepared by ML. GOCART model experiments were prepared by TK. HadGEM2-ES model experiments were prepared by GF. OsloCTM3.v2 model experiments were prepared by MTL. RAQMS model experiments were prepared by RBP and AL. SPRINTARS model experiments were prepared by TT. UI provided AQMEI1I3 regional ensemble model output for comparison. FJD and TJK coordinated HTAP2. All authors commented on drafts of the paper.

Competing interests. The authors declare that they have no conflict of interest.

References


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